

Function of a random variable

Let X be a random variable in a probabilistic space (Ω, \mathcal{S}, P) with a probability distribution $F(x)$

Sometimes we may be interested in another random variable Y that is a function of X , that is, $Y = g(X)$. The question is whether we can establish the probability distribution $G(y)$ of Y .

Let $g(x)$ be increasing on $R(X)$.

Since $F(x)$ is the probability distribution of X and $G(y)$ of Y , we can write

$$F(a) = P(X < a) \quad \text{and} \quad G(b) = P(Y < b) = P(g(X) < b)$$

$g(x)$ is increasing and so it has an inverse $g^{-1}(x)$

$$P(g(X) < b) = P(g^{-1}(g(X)) < g^{-1}(b)) = P(X < g^{-1}(b)) = F(g^{-1}(b))$$

This means that $G(y) = F(g^{-1}(y))$

If $g(x)$ is decreasing on $R(X)$, we can proceed in a similar way, but since the inverse of a decreasing function is again decreasing, the inequality will revert the relation sign and so

$$\begin{aligned} P(g(X) < b) &= P(g^{-1}(g(X)) \geq g^{-1}(b)) = \\ &= P(X \geq g^{-1}(b)) = 1 - P(X < g^{-1}(b)) = 1 - F(g^{-1}(b)) \end{aligned}$$

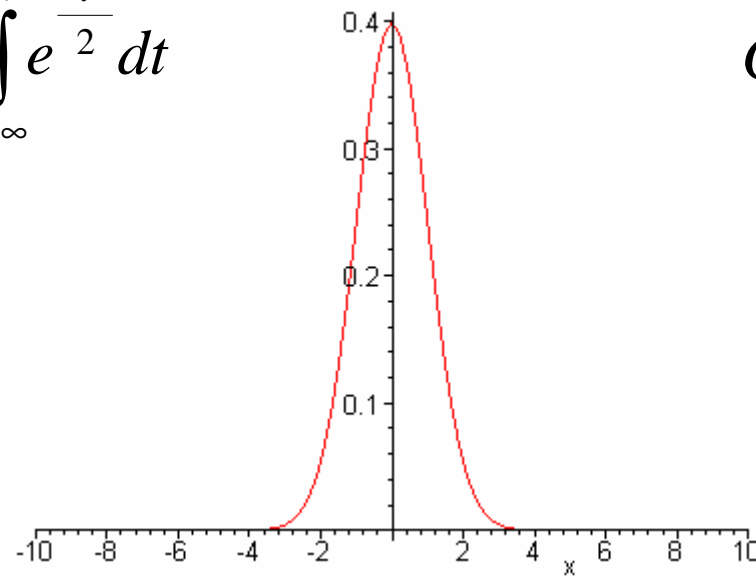
$$\text{so that } G(y) = 1 - F(g^{-1}(y))$$

Sometimes we can use even more sophisticated methods as shown in the following example.

Random variable X has the standardized normal distribution $\Phi(x)$, calculate the distribution of the random variable $Y = X^2$.

$$\Phi(x) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^x e^{-\frac{t^2}{2}} dt$$

$$G(y) = ?$$



$$G(y) = P(X^2 < y) = P(-\sqrt{y} < X < \sqrt{y}) = \Phi(\sqrt{y}) - \Phi(-\sqrt{y})$$

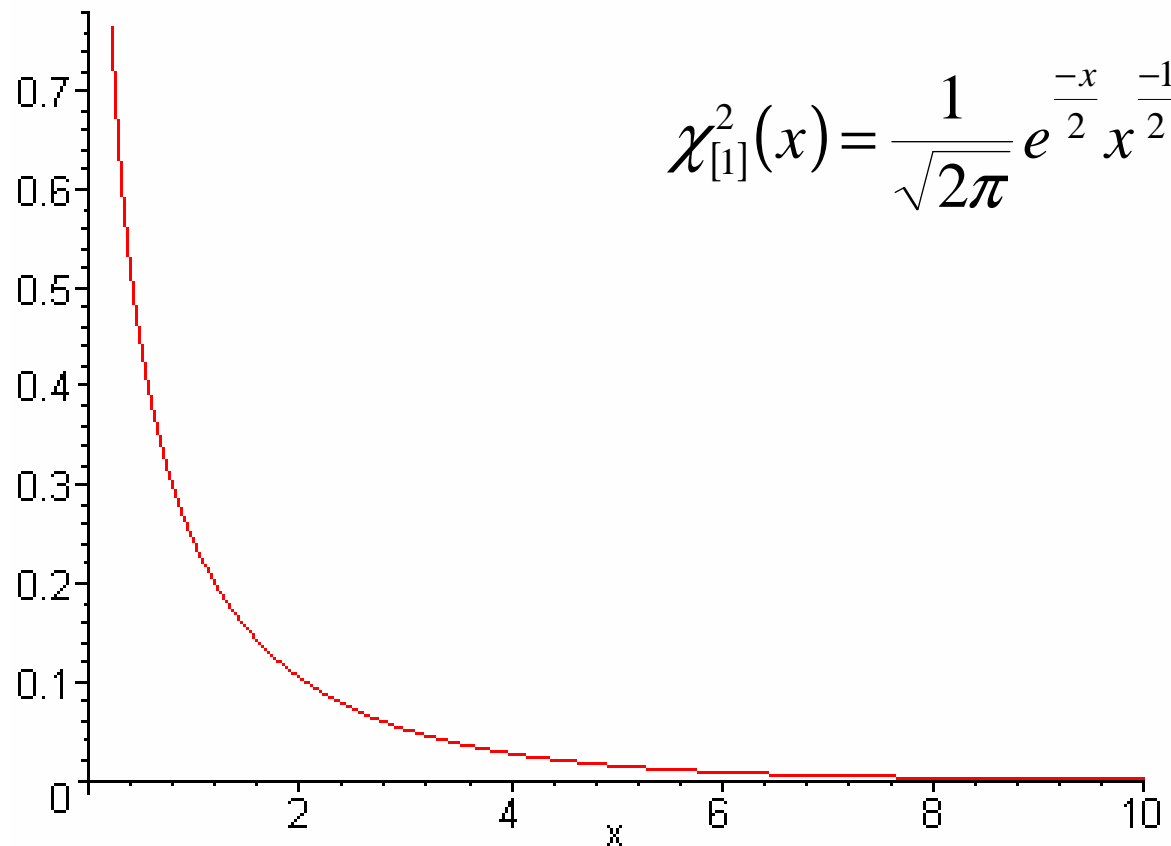
$$\Phi(\sqrt{y}) - \Phi(-\sqrt{y}) = \Phi(\sqrt{y}) - (1 - \Phi(\sqrt{y})) = 2\Phi(\sqrt{y}) - 1$$

$$G(y) = \frac{2}{\sqrt{2\pi}} \int_{-\infty}^{\sqrt{y}} e^{-\frac{t^2}{2}} dt - 1$$

We cannot calculate the last integral exactly, but we can establish the probability density of Y by differentiating $G(y)$

$$g(y) = \frac{dG(y)}{dy} = \frac{2}{\sqrt{2\pi}} e^{-\frac{y}{2}} \frac{1}{2\sqrt{y}} = \frac{1}{\sqrt{2\pi}} e^{-\frac{y}{2}} y^{-\frac{1}{2}}$$

This is the probability density of the random variable X^2 . We have, in fact, calculated the density of the chi-squared distribution with one degree of freedom.



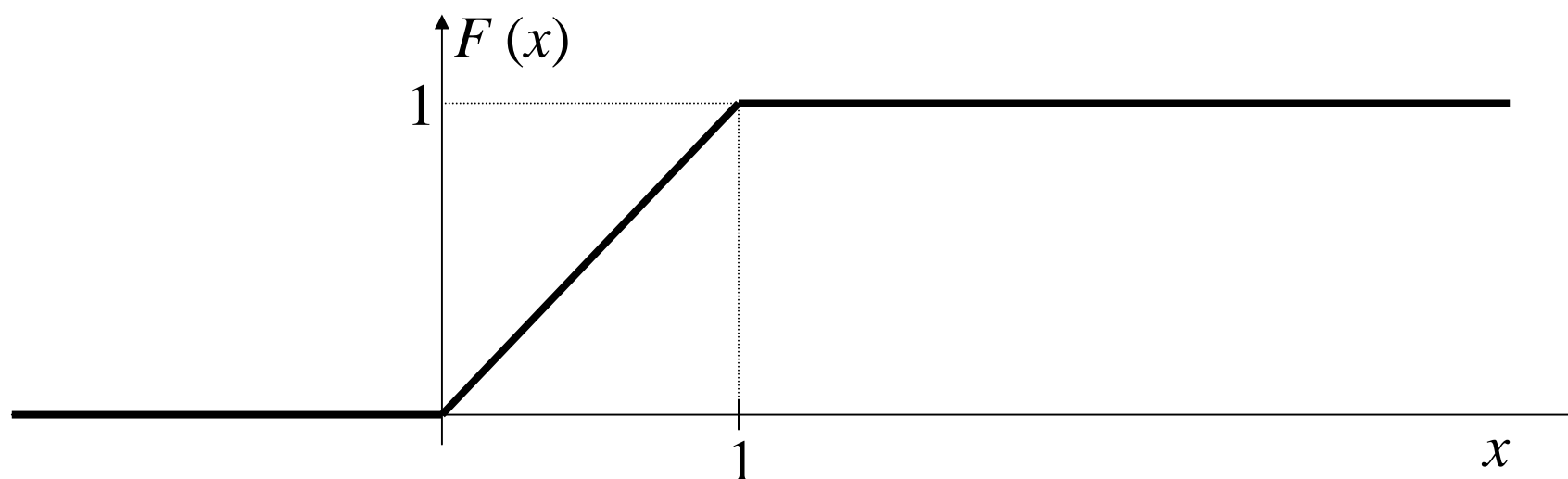
Example

Random variable X has a uniform distribution on $[0,1]$. Find a transformation $g(X)$ such that $Y = g(X)$ has a distribution $F(y)$.

X has the density $f(x) = 1$ for $x \in [0,1)$, $f(x) = 0$ otherwise

and the distribution

$$F(x) = 0 \text{ for } x < 0, F(x) = x \text{ for } x \in [0,1), F(x) = 1 \text{ for } x \geq 1$$



$F(x)$ is a distribution and as such has R as the domain and $[0,1]$ as the range. This means that, if $F(x)$ is increasing, it has an inverse $F^{-1}(x)$ that is also increasing with range R and domain $[0,1]$.

Consider the transformation $Y = F^{-1}(x)$. We have

$$\begin{aligned} P(Y < a) &= P(F^{-1}(X) < a) = P(F(F^{-1}(X)) < F(a)) = \\ &= P(X < F(a)) = F(a) \end{aligned}$$

Thus, we have shown that $F(y)$ is the distribution of $Y = F^{-1}(x)$.

This can be used for example for simulating a distribution using a pocket calculator.